A Method for Directly Generating a Gaussian Distribution with Nonunit Variance and Nonzero **Mean from Uniform Random Deviates**

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- 1. Introduction

variance: $G(x) = \frac{e}{\sqrt{2\pi}}$. For nonzero mean μ and nonunit variance σ^2 , the Gaussian distribution $\left(-\frac{(x-\mu)^2}{2}\right)$ $\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$ is $G(x, \mu, \sigma) = \frac{e}{\sigma\sqrt{2\pi}}$, with normalization $\int_{0}^{\infty} G(x, \mu, \sigma) dx = 1$. How does one obtain a

distribution with a given variance from a distribution with unit variance? The Gaussian function has a useful invariant relation: $G(x, \mu, \sigma) = \frac{G(y, 0, 1)}{\sigma}$, where $y = \frac{x - \mu}{\sigma}$. Hence, if we have a set of random deviates y which are distributed as a Gaussian with unit variance and zero mean, we may convert those deviates to a Gaussian distribution with nonzero mean and nonunit variance by the

simple trick of multiplying by σ and adding μ . However, one might still wish for an algorithm that takes uniform deviates and *directly* produces the desired nonunit variance Gaussian distribution.

2. A Modified Box-Muller Algorithm

The usual algorithm for producing a Gaussian distribution from uniform random deviates makes use of the Box-Muller transformation. We will modify the usual algorithm so that the Box-Mueller transformation produces the desired nonunit-variance Gaussian distribution. If z_1 and z_2 are independent and uniformly distributed in the range 0 to 1, then consider y_1 and y_2 , given by

$$y_1 = \sqrt{-2 \ln(z_1)} \cos(2 \pi z_2)$$
$$y_2 = \sqrt{-2 \ln(z_1)} \sin(2 \pi z_2)$$

This is the Box-Muller transformation, which converts a 2D uniform distribution (z_1, z_2) to a bivariate Gaussian distribution (y_1, y_2) with zero mean and unit variance. (The proof that this transformation yields the desired Gaussian distribution is a special case of the proof of the modified transformation, which is given below.)

The trick to making the Box-Muller transformation produce a Gaussian distribution (x_1, x_2) with nonunit variance is to use the invariant relation mentioned above. Setting

$$x = v \sigma + \mu$$

we have

$$x_1 = \sqrt{-2 \ln(z_1)} \cos(2 \pi z_2) \sigma + \mu$$

 $x_2 = \sqrt{-2 \ln(z_1)} \sin(2 \pi z_2) \sigma + \mu$

2.1. Verification

To show that this is indeed the desired distribution, we have

$$\sqrt{\frac{(x_1 - \mu)^2}{\sigma^2} + \frac{(x_2 - \mu)^2}{\sigma^2}} = \sqrt{2} \sqrt{-\ln(z_1)}$$

which we solve for z_1 ,

$$z_{1} = \mathbf{e}^{\left(-\frac{1}{2} \frac{(x_{2} - \mu)^{2}}{\sigma^{2}}\right)} \mathbf{e}^{\left(-\frac{1}{2} \frac{(x_{1} - \mu)^{2}}{\sigma^{2}}\right)}$$

and we have

$$\frac{x_2}{x_1} = \frac{\sqrt{-2 \ln(z_1)} \sin(2 \pi z_2) \sigma + \mu}{\sqrt{-2 \ln(z_1)} \cos(2 \pi z_2) \sigma + \mu}$$

which we solve for z_2 ,

$$\begin{split} z_2 &= -\frac{1}{2} \arctan \left(\left(\left(\sqrt{-\ln(z_1)} \ x_1 \ (-x_1 + x_2) \ \mu \right) \right. \\ &- \sqrt{\ln(z_1) \ x_2^2 \ (-x_1 + x_2)^2 \ \mu^2 + 2 \ln(z_1)^2 \ x_2^2 \left(x_2^2 + x_1^2 \right) \sigma^2} \ \right) x_2 \right) \bigg/ \left(\\ &\sqrt{-\ln(z_1)} \ x_2^2 \ (-x_1 + x_2) \ \mu \\ &+ x_1 \sqrt{\ln(z_1) \ x_2^2 \ (-x_1 + x_2)^2 \ \mu^2 + 2 \ln(z_1)^2 \ x_2^2 \left(x_2^2 + x_1^2 \right) \sigma^2} \ \right) \bigg) / \pi \end{split}$$

Substituting for z_1 and simplifying, we find

$$z_2 = \frac{1}{2} \frac{\arctan\left(\frac{-x_2 + \mu}{-x_1 + \mu}\right)}{\pi}$$

Now, the fundamental transformation law of probabilities states that |p(y)| dy = |p(x)| dx, or $p(y) = p(x) \left| \frac{dx}{dy} \right|$, where p(x)| dx is the probability that x lies between x and x + dx, and p(y)| dy is the probability that y lies between y and y + dy. The extension of this to n dimensions is p(Y)| dY = p(X)| J(X, Y)| dY, where p(X) is the joint probability distribution of Y, $Y = (x_1, x_2 ... x_n)$, $Y = (y_1, y_2 ... y_n)$, $dY = dy_1| dy_2 ... dy_n$, and J(X, Y)

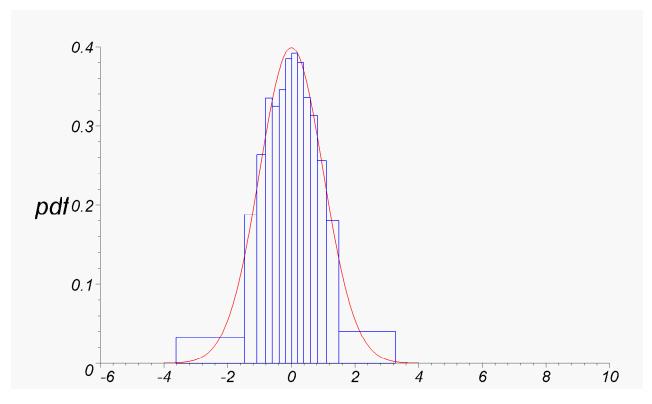
is the Jacobi determinant. For two dimensions, $J(X, Y) = \det \begin{bmatrix} \frac{\partial}{\partial y_1} x_1 & \frac{\partial}{\partial y_2} x_1 \\ \frac{\partial}{\partial y_1} x_2 & \frac{\partial}{\partial y_2} x_2 \end{bmatrix}$. Hence,

$$J(Z, X) = -\frac{1}{2} \frac{e^{\left(-x_1 + \mu\right)^2}}{\sigma^2} \int_{e}^{\left(-1/2 + \mu\right)^2} \frac{\left(-x_2 + \mu\right)^2}{\sigma^2}$$

This says that for a uniform distribution (z_1, z_2) the distribution of (x_1, x_2) is a symmetric bivariate Gaussian distribution with mean μ and variance σ^2 . *Q.E.D.*

3. Numerical Tests

In the plot below, the red curve is a plot of G(x, 0, 1), while the histogram is the result of running 5000 uniform random deviates through the modified Box-Muller transformation just described with $\mu = 0$ and $\sigma = 1$.

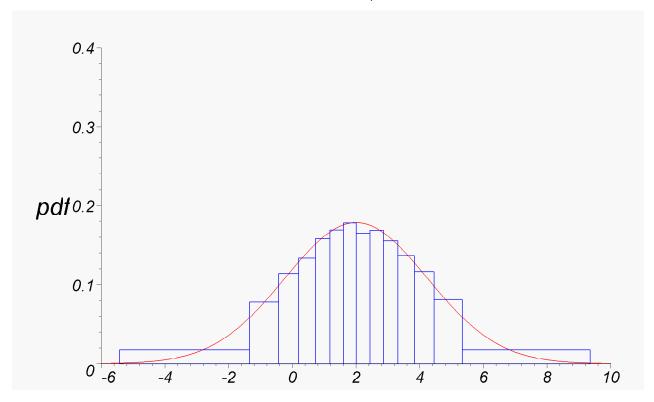


The peak of the curve should be

$$\frac{1}{\sqrt{2 \pi}} = .3989422802$$

This shows that the modified numerical algorithm correctly reproduces the standard algorithm in the appropriate limits.

Now for a test of the modified algorithm. In the next plot, the red curve is a plot of $G(x, 2, \sqrt{5})$, while the histogram is the result of running 5000 uniform random deviates through the modified Box-Muller transformation with $\mu = 2$ and $\sigma = \sqrt{5}$.



The peak of the curve should be

$$\frac{1}{\sqrt{5}\sqrt{2\pi}} = .1784124116$$

This shows that the modified algorithm is indeed generating the correct Gaussian distribution with nonunit variance and nonzero mean.